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Uncertainty in model projections of future climate change arises due to internal variability, multiple possible emission scenarios, and different model responses to anthropogenic forcing. To robustly quantify uncertainty in multi-model ensembles, inter-dependencies between models as well as a model's ability to reproduce observations should be considered. Here, a model weighting approach, which accounts for both independence and performance, is applied to European temperature and precipitation projections from the CMIP5 archive. Two future periods representing mid- and end-of-century conditions driven by the high-emission scenario RCP8.5 are investigated. To inform the weighting, six diagnostics based on three observational estimates are used to also account for uncertainty in the observational record. Our findings show that weighting the ensemble can reduce the interquartile spread by more than 20% in some regions, increasing the reliability of projected changes. The mean temperature change is most notably impacted by the weighting in the Mediterranean, where it is found to be 0.35 °C higher than the unweighted mean in the end-of-century period. For precipitation the largest differences are found for Northern Europe, with a relative decrease in precipitation of 2.4% and 3.4% for the two future periods compared to the unweighted case. Based on a perfect model test, it is found that weighting the ensemble leads to an increase in the investigated skill score for temperature and precipitation while minimizing the probability of overfitting.

1. Introduction

Changes in the climate system due to anthropogenic greenhouse gas emissions affect society through rising temperatures and changing precipitation patterns, among other things (IPCC 2013, WMO 2019). Efforts are ongoing to limit global mean warming to well below 2 °C in order to avoid dangerous levels of climate change (United Nations 2015). On a regional scale, however, temperature and precipitation changes can vastly differ from the global average (Seneviratne *et al* 2018) and uncertainties are often larger (Hawkins and Sutton 2009, Deser *et al* 2012a). On such scales it is, therefore, even more important for decision makers, such as politicians and large infrastructure providers, to make decisions based on the best available knowledge.

Efforts are ongoing in the scientific community to deliver such reliable and actionable information on relevant temporal and spatial scales, for example, within the European Climate Prediction system (EUCP) project (Hewitt and Lowe 2018). The EUCP aims to combine regional and global climate model projections from the latest Coupled Model Inter-comparison Projects CMIP5 (Taylor *et al* 2012) and CMIP6 (Eyring *et al* 2016) with targeted statistical approaches to reduce uncertainty in multi-model mean projections of regional climate.

Model projections are affected by a range of uncertainties including emissions scenario uncertainty, internal variability of the climate system, and model response uncertainty (Hawkins and Sutton 2009). Scenario uncertainty takes into account different possible socio-economic developments (van Vuuren *et al* 2011,

O'Neill *et al* 2014) and is as such mainly determined by political decisions and technological development. Internal variability arises from the chaotic behavior of the weather and can not be reduced when investigating a given region and time period of interest (Hawkins and Sutton 2009, Deser *et al* 2012b, Fatichi *et al* 2016).

Hence, reducing the model spread is often the only option to reduce total uncertainty of climate model projections. Many different approaches have been used in the past, based on the assumption that the reliability of a model in the future can be inferred from its performance in reproducing historical climate. These approaches argue that some models may be more 'fit for purpose' for a given application (e.g. variable, region, or lead time) than others (Gleckler *et al* 2008, Parker 2009, Knutti 2010, Baumberger *et al* 2017, Eyring *et al* 2019), and thus deserve to receive more weight in the multi-model estimate of change.

In addition, when working with an 'ensemble of opportunity' (Tebaldi and Knutti 2007) such as CMIP5, the handling of model inter-dependencies becomes an important question to address. The models contributing to CMIP5 were not designed to be independent from each other: developers might have shared ideas and code, some models might use identical components, or two models might have been branched from the same parent (Knutti *et al* 2013, Sanderson *et al* 2015a, Annan and Hargreaves 2017, Abramowitz *et al* 2019, Eyring *et al* 2019). Moreover, for several models, multiple realizations with slightly perturbed initial conditions are provided. On the one hand, such ensembles should not simply be included in a multi-model estimate, since they represent the same model, hence artificially increasing the weight of models with more members. On the other hand, using only one member per model discards the additional information provided by these initial condition ensembles.

The raw model range of the full CMIP5 ensemble, hence, does not necessarily represent the actual uncertainty of the climate system. It rather reflects the ensemble design (or a lack thereof) and can either overestimate or underestimate the actual uncertainty (Lorenz *et al* 2018). To better capture real uncertainty different methods account for model performance and independence by selecting only a subset of models (Pennell and Reichler 2011, Evans *et al* 2013, Leduc *et al* 2016, Hegerl *et al* 2018, Hegerl *et al* 2019), finding emergent constraints for the change in an investigated variable (Hall and Qu 2006, Gordon and Klein 2014, Cox *et al* 2018, Hall *et al* 2019), or model weighting (Coppola *et al* 2010, Abramowitz and Bishop 2015, Knutti *et al* 2017, Sanderson *et al* 2017, Lorenz *et al* 2018).

We here use a model weighting approach which accounts for model performance and independence (Sanderson *et al* 2015a, 2015b, Knutti *et al* 2017, Lorenz *et al* 2018). In an initial step our study focuses on selecting relevant diagnostics to establish model

performance and independence. For the first time we use multiple observational data sets to inform this method and to include an estimate of observational uncertainty. Calculated weights are then applied to projections of temperature and precipitation change in eight different regions in Europe and in two future periods. The selection of these regions and time periods is partly guided by a coordinated effort by the EUCP to compare different methods constraining European climate projections to which this work is contributing. In addition, the selected regions allow for a discussion of the weighting performance in different parts of Europe and for three distinctively different regions sizes. Ultimately the aim of this study is to better quantify the uncertainty of future changes and to improve projection skill as well as enable the usage of full multi-model ensembles without having to manually account for multiple initial condition members or model dependencies.

2. Data and methods

2.1. Model data

This study is based on all available CMIP5 models which contain the variables required to inform the weighting (37 models with a total of 79 realizations; see figure S6 in the supplement for a full list). We use monthly data, regridded to a regular $2.5^\circ \times 2.5^\circ$ grid using bi-linear remapping. Throughout the study, we base our weighting approach on the base-period 1995–2014 in the combined historical and Representative Concentration Pathway 8.5 (RCP8.5) (van Vuuren *et al* 2011) forcing runs, and apply the weights to RCP8.5 forcing runs for the future projections. The base-period is selected to (i) best represent present day conditions and (ii) in anticipation of and to be easily comparable to the upcoming CMIP6 which uses historical forcing up to 2014 (Eyring *et al* 2016).

2.2. Observation-based datasets

To account for uncertainties in the observational records, the performance weighting is based on observational estimates from two reanalysis datasets and one combined observational dataset. We use the interim reanalysis from the European Centre for Medium-Range Weather Forecasts (ERA-Interim) (Dee *et al* 2011) and the Modern-Era Retrospective analysis for Research and Applications version 2 from the National Aeronautics and Space Administration (MERRA2) (Gelaro *et al* 2017). In addition, we use a combined observational dataset, which is based on the ensemble version 17 of E-OBS (Cornes *et al* 2018) (for temperature, precipitation, sea level pressure) and the CERES EBAF Surface Ed2.8 (Kato *et al* 2013) (for radiation variables).

Table 1. Regions and shape parameters (σ_D and σ_S) used in this study. The coordinates for the last four regions give the center of the $2.5^\circ \times 2.5^\circ$ grid cell used.

Key	Name	Location	σ_D (tas)	σ_D (pr)	σ_S
EUR	Europe	NEU + CEU + MED	0.588	0.658	0.704
NEU	Northern Europe	NEU	0.614	0.542	0.735
CEU	Central Europe	CEU	0.607	0.831	0.706
MED	Mediterranean	MED	0.546	0.667	0.643
FAL	Falun (Sweden)	13.75 °E, 61.25 °N (NEU)	0.852	0.800	0.628
DUS	Dusseldorf (Germany)	6.25 °E, 51.25 °N (CEU)	0.929	0.956	0.790
SIB	Sibiu (Romania)	23.75 °E, 46.25 °N (CEU)	1.030	1.227	0.691
MAD	Madrid (Spain)	-3.75 °E, 41.25 °N (MED)	1.116	1.089	0.604

Table 2. Variables tested as diagnostics as well as selected diagnostics. Possible time aggregations are climatology (CLIM) and standard deviation (STD).

Key	Name	Mask	Used diagnostics
pr	Precipitation	Sea	pr-CLIM
psl	Sea level pressure	—	
rnet	Surface downwelling net radiation	Sea	
rlds	Surface downwelling longwave radiation	Sea	rlds-STD
rlus	Surface upwelling longwave radiation	Sea	
rsds	Surface downwelling shortwave radiation	Sea	rsds-CLIM
rsus	Surface upwelling shortwave radiation	Sea	rsus-CLIM, rsus-STD
tas	Near-surface air temperature	Sea	tas-CLIM
tos	Sea surface temperature	Land	

2.3. Model weighting

We focus on changes in summer (JJA) temperature and precipitation in the three European SREX (IPCC 2012) regions: Northern Europe (NEU), Central Europe (CEU), and the Mediterranean (MED) as well as in the combined European domain (EUR). In addition, we test the performance of the method on four single grid cells as listed in table 1. A geographical map showing all regions as well as the model resolution and the applied land-sea mask can be found in the supplement (figure S1 is available online at stacks.iop.org/ERL/14/124010/mmedia). The effects of the weighting on future projections for two 20-year periods representing mid-century (2041–2060) and end-of-century (2081–2100) conditions are investigated.

Weights are calculated for each model following the approach presented by Lorenz *et al* (2018), which is based on earlier work from Knutti *et al* (2017) and Sanderson *et al* (2015a, 2015b). Each weight w_i is a combination of the observational distance D_i (informing the performance weighting) and the model distance S_{ij} (informing the independence weighting):

$$w_i = \frac{e^{-\frac{D_i}{\sigma_D}}}{1 + \sum_{j \neq i}^M e^{-\frac{S_{ij}}{\sigma_S}}}, \quad (1)$$

with the total number of model runs M and the shape parameters σ_D and σ_S . The shape parameters define the strength of the weighting and the relative importance of performance and independence, large values will lead to an approximation of equal weighting, while small values will lead to aggressive weighting, giving a few models most of the weight. The shape parameters

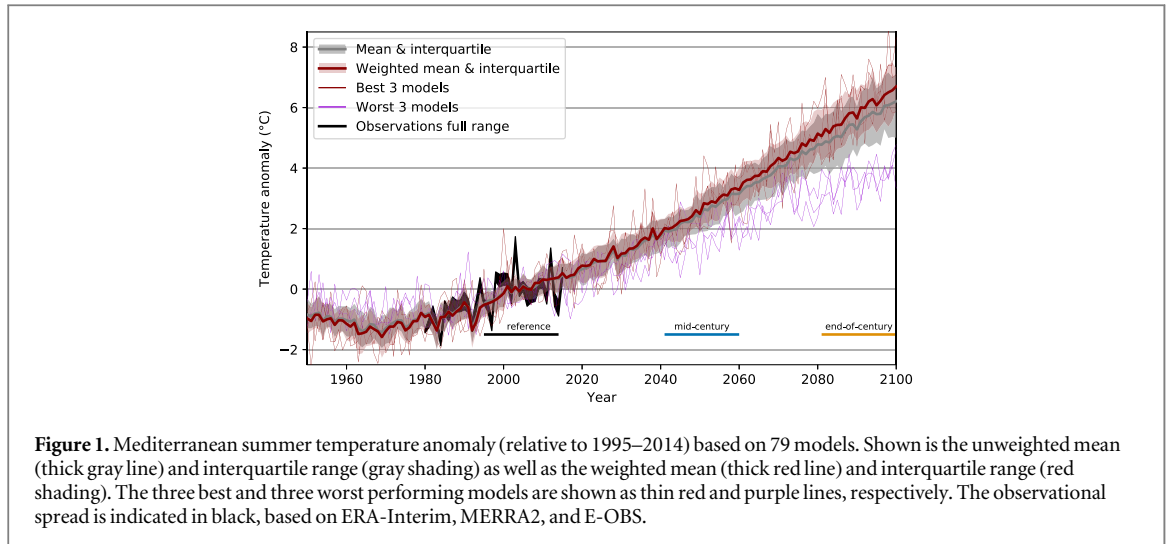
used in this study are summarized in table 1 and their calculation is described by Lorenz *et al* (2018) and in the supplement.

2.4. Diagnostics

The observational distance D_i is based on six diagnostics (see table 2). A diagnostic can be based on any CMIP5 output variable for which observations are available. The calculation of D_i follows a straightforward approach: (i) variable, region, time period, and season are selected, (ii) the mean (CLIM) or standard deviation (STD) over the time period is calculated, (iii) the point-to-point distance between model and the center of the observational spread is calculated, and finally (iv) the area-weighted root mean squared error is calculated over the selected region.

To account for uncertainty in our knowledge of observed climate we use three observational estimates in the calculation of the point-to-point distance (iii). Several methods of combining these estimates and their influence on the weighting are investigated. Distances are tested with regard to each of the estimates individually as well as with regard to mean, median, and full spread. A detailed description of all steps and the different methods of combining observations can be found in the supplement.

The independence weighting is informed by diagnostics which can be based on any CMIP5 output variable which is available for all models. In practice, we use the same diagnostics as for the performance weighting. To compute the model distance S_{ij} , the



point-to-point difference (step (iii) above) is calculated between each model pair.

To select the most informative diagnostics, different combinations and numbers of diagnostics are tested and evaluated. The aim is to select diagnostics which are, on the one hand, relevant for the target and, on the other hand, add additional information to the already used diagnostics (Gleckler *et al* 2008, Lorenz *et al* 2018). We test a pool of variables (table 2) which we judge to be relevant for predicting our targets and for which observations are available. Indeed, the availability of observations of sufficient duration, spatial coverage, and quality is a considerable constraint for the pool of tested variables, particularly since we here aim at using multiple data sets in order to include observational uncertainty. Using the same approach as Lorenz *et al* (2018), we select six diagnostics, which is found to be a good number to differentiate between high and low performing models (performance tends to converge with an increasing number of diagnostics) while not being overconfident in a model's ability to reproduce observed climate (which becomes an issue with a low number of diagnostics) (Lorenz *et al* 2018).

2.5. Model-as-truth testing

A model-as-truth test (or perfect model test) picks each model from a multi-model ensemble in turn and treats it as the true representation of the climate system. Weights for the remaining models are computed using historical information from this perfect model as 'pseudo-observations'. This allows for an evaluation of the impact of the weighting in the future based on each model representing the truth once. Here this approach is used for two applications: (i) to estimate the ideal shape parameter for the performance weighting σ_D (as detailed in the supplement) and (ii) to evaluate the impact of the weighting on the skill of the future projections. As a measure for the skill we use the continuous ranked probability score (CRPS) as detailed in Hersbach (2000). For each perfect model the CRPS basically represents the mean

absolute error between the distribution of all other models and the perfect model. We define the skill of the weighting as the relative change in CRPS in the end-of-century period:

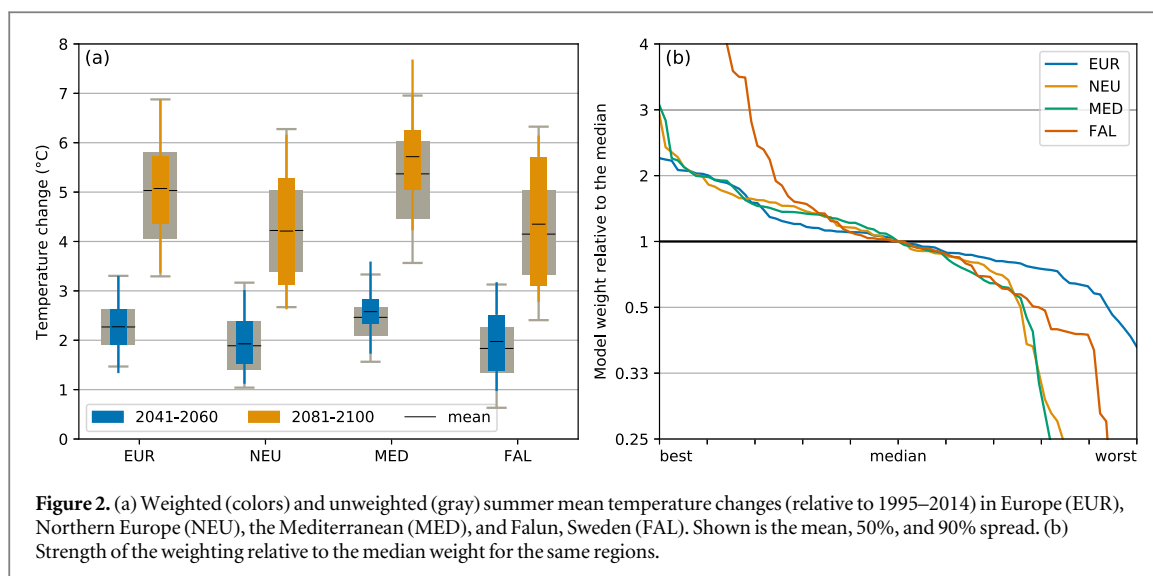
$$\text{skill} = \frac{\text{CRPS}_{\text{unweighted}} - \text{CRPS}_{\text{weighted}}}{\text{CRPS}_{\text{unweighted}}} \times 100\%. \quad (2)$$

3. Results

3.1. Summer temperature changes

We first calculate the optimized shape parameters (σ_D and σ_S) for all regions and periods. They deviate only slightly between mid-century and end-of-century and we therefore use the average over both periods in the computation of weights (table 1). Across all three SREX regions, the sigma values for temperature average to about $\sigma_D = 0.6$ and $\sigma_S = 0.7$, which is comparable to earlier work focusing on maximum summer temperature in North America (Lorenz *et al* 2018). In the four grid cell regions, the performance sigma values are about 50% higher than for the SREX regions due to the larger range of model-observation differences without spatial aggregation. Generally speaking, weighting based on smaller regions is less robust and, therefore, more likely to be overconfident. As a consequence the perfect model test correctly picks higher performance sigma values that lead to more evenly distributed weights and reduce overfitting.

Weighted and unweighted time-series of Mediterranean summer temperature are shown in figure 1. Models with less warming consistently get the lowest weights and, as a consequence, the weighted multi-model mean shifts upward by an additional 0.35°C by the end of the century. The 25th percentile also shows a strong increase, moving above 5°C warming for 2081–2100 (see also figure 2(a)). Since the increase in the 75th percentile is less distinctive, the weighting



also leads to a reduction of the interquartile spread by almost 25%.

Figure 2(a) shows temperature change distributions for the combined European domain, two SREX regions, and one of the grid cell regions. A summary of all regions can be found in figure S2(a) in the supplement. The effect of the weighting is generally smaller for the mid-century period. In the full European domain there is hardly any shift in the mean change, but a considerably reduction of the interquartile spread by about one quarter for the end-of-century period. Analyzing spatial differences reveals that the European domain seems to be a too large region to average over. The weighted mean is colder in the northern part and warmer in the southern part compared to the unweighted case consistent with the SREX regions described below. Aggregating over the entire domain cancels out most of this effect in the mean. In Northern Europe the interquartile spread widens by almost a third due to the weighting. This behavior is even stronger for the Falun grid cell with an increase by over 50%. Based on our diagnostics, it seems that models on both ends of the future temperature distribution get higher weights in NEU, indicating that a simple average and standard deviation might be slightly overconfident.

In Central Europe (figure S2(a)) neither mean nor spread are strongly affected by the weighting. This somewhat differs from findings by Vogel *et al* (2018), who constrain summer maximum temperature in this region to well below the ensemble mean, using the correlation between maximum temperature and precipitation as diagnostic. However, work from Stegehuis *et al* (2013) suggests that warming is underestimated by models in Northern and Central Europe and overestimated in the Mediterranean. We argue, therefore, that our approach might give a more robust estimate of the potential to narrow the uncertainty in this region, since it is based on a range of

six diagnostics rather than on a single emergent constraint. It is, hence, not surprising that the Central European temperature distribution is not strongly affected by the weighting in our more conservative framework.

For the Mediterranean the weighting suggests that the unweighted multi-model mean strongly underestimates the potential warming in both periods. According to the weighted distribution, temperature increases of more than 2 °C are very likely already by the middle of the century when following RCP8.5. For the end of the century warming of over 5 °C compared to the 1995–2014 reference becomes increasingly likely.

Figure 2(b) shows that the best models receive up to tree times the median weight, except for the small grid cell regions Falun and Madrid with a maximum of 14× and 7× higher weights, respectively (see also figure S2(b)). The worst performing models in turn can get considerably lower weights more than 20× smaller than the median in several cases. Combining the results from figures 2(a) and (b) reveals that the effect of the weighting on the temperature distribution is a combination of the weighting strength and whether weights are correlated with simulated warming. The strong shift in the temperature distribution in the Mediterranean region is mainly driven by about 20% of the models which are heavily down-weighted for being far away from the observations in the historical period. For this specific region and target the majority of these models come from the GISS model family, which projects temperature changes on the lower end of the distribution. In Central Europe the weighting is equally strong but there is no clear clustering of weights for warmer or colder models so that there is hardly any effect on the temperature distribution. A list of all models included in the study and their corresponding weights in each of the regions can be found in figure S6 in the supplement.

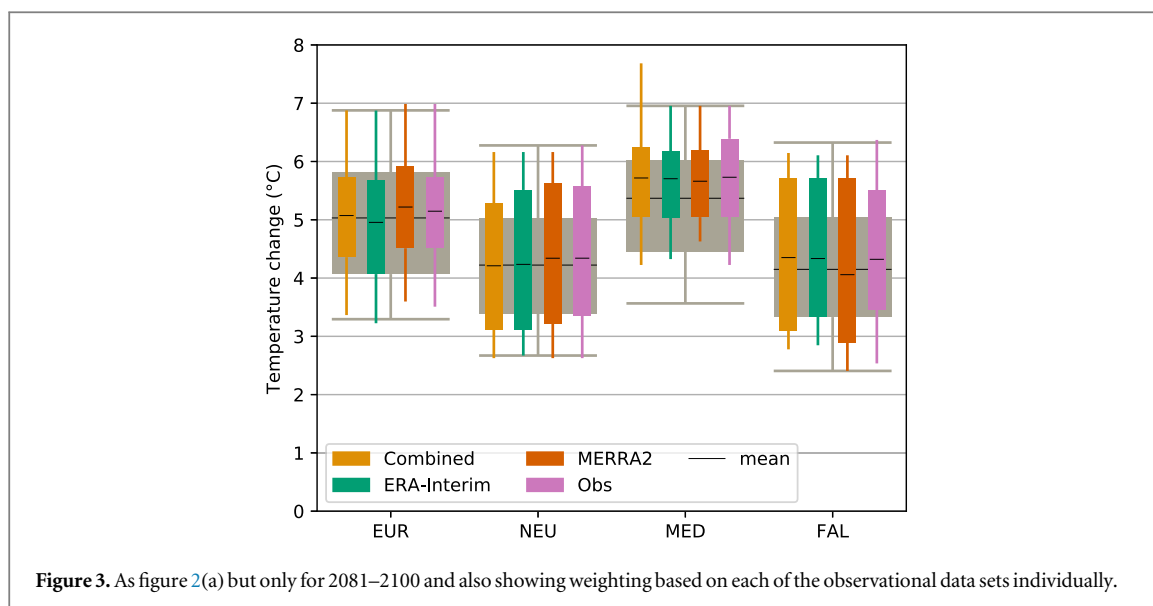


Figure 3. As figure 2(a) but only for 2081–2100 and also showing weighting based on each of the observational data sets individually.

3.2. The effect of observational uncertainty

The area-mean observational range for the Mediterranean (shown in black in figure 1) can cover large fractions of the interquartile model spread. This highlights the importance of accounting for observational uncertainty, particularly for diagnostics where many models lie within the observational spread. To investigate the effect of using multiple observational estimates we also calculate the weights based on each of the three used observational datasets separately. In most regions using only one dataset can lead to a shift in the mean which is not in line with the combined case (figure 3). This becomes even more important when considering smaller regions such as FAL, where using different observational estimates can lead to the weighting changes having different signs. Note that combining multiple datasets does not necessarily lead to a linear combination of the weighting effect. This can best be seen in the mean of NEU or in the 95th percentile of MED. We therefore stress the importance of using as many observational datasets as possible or feasible to best capture uncertainty in the observational estimates when evaluating historical model performance.

To investigate the robustness of our approach to combine multiple observational datasets we also look into additional methods of combining the information from the three datasets (see supplement for a detailed description). As discussed in detail in section 4 one can even argue that the observations should also be weighted, but since the results from the different approaches tested are reasonably robust (figure S3), we limit ourselves to using the center of the observational spread here.

3.3. Summer precipitation changes

Precipitation changes are most notably affected by the weighting in Northern Europe (figure 4(a)): future precipitation increases are strongly constrained for

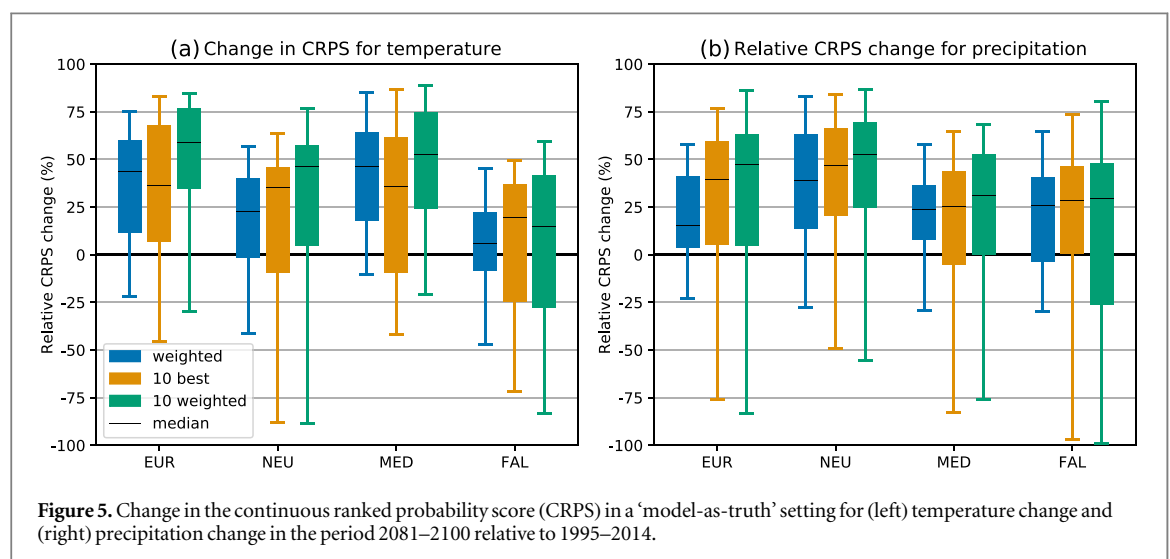
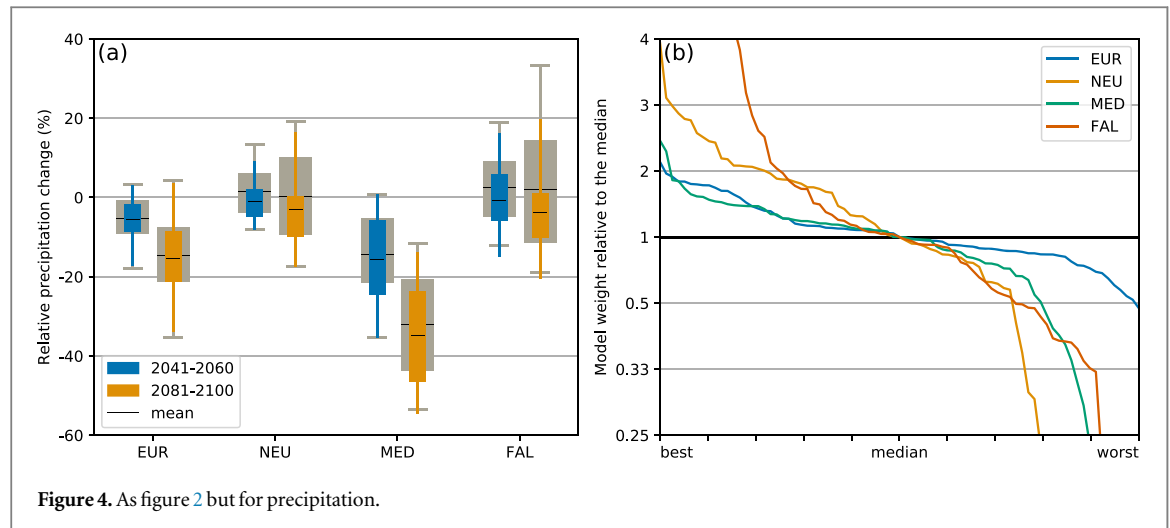
both periods. This drives a reduction of the interquartile range by up to 50%. Similar to temperature, the effect on the precipitation distribution is mainly driven by a strong down-weighting of about 25% of models mainly from the CSIRO and IPSL families, which show a slight precipitation increase (figures 4(b) and S6). Looking into the absolute precipitation rather than the change reveals that the lower weighted models are predominantly drier than the mean in the historical period and subsequently show a slightly stronger increase in precipitation than the mean under RCP8.5.

In the combined European domain as well as Central Europe, the weighting leads to a reduction in the uncertainty of the projections of about 20% in the middle of the century as the most prominent effect. Using precipitation observations as a constraint this potential of significantly reducing uncertainty in future precipitation changes was also found by, e.g. Zhang and Soden (2019). In the Mediterranean the weighted distribution indicates a tendency towards even dryer conditions than the unweighted mean. Together with the results from temperature shown in figure 2(a) this suggests that models with higher temperatures and associated stronger drying are more reliable based on their past performance.

3.4. Projection skill changes in model-as-truth tests

To evaluate the skill of the weighting approach we use the CRPS in model-as-truth experiments for the end of the century (2081–2100). Figure 5 shows the skill for three different cases: (i) weighing of the full ensemble, (ii) selecting a subset of the 10 best models (based on the same weights) with each model getting equal weight, and (iii) selecting a subset of the 10 best models where each models keeps its original weight.

For temperature in the European domain the median skill increases by almost 50% due to the weighting. In the three SREX regions the skill also



increases, while there is hardly any increase in the grid cell region (figures 5(a) and S5(a)). It is noteworthy that the skill increase is also observable in regions where there is no obvious change in the temperature distribution (see, e.g. CEU in figure 2(a)). In the 'models are indistinguishable from the truth' paradigm (Sanderson and Knutti 2012) this can be interpreted as the observations representing one realization which does not lead to a strong weighting. But other realizations (represented by model 'pseudo-observations') can lead to a weighting which yields a skill increase.

For all regions there is also a chance that the skill decreases due to the weighting. This can happen if the perfect model has a very different response to future forcing compared to the other models, leading to the weighted multi-model ensemble moving further away from the 'truth'. This represents the possibility that weighting models based on the distance to the observations may have negative impact on the skill of the future projections. To minimize this risk of overconfidently weighting the projections and giving too narrow uncertainty estimates we select the performance shape parameter σ_D also based on a model-as-truth test.

Therefore, the probability for a decrease in the skill is small when using the full ensemble.

Comparing this approach to picking only the 10 best models (unweighted or weighted), yields similar results for the median CRPS change and for the inter-quartile range. For most regions the median performance is even slightly better than for the full weighted ensemble. Crucially, however, picking only 10 members can perform considerably worse in the lower percentiles. Selecting a subset of models has a higher chance of overfitting and can lead to a reduction in the CRPS by more than 75% for temperature.

Depending on the question it might, therefore, be sufficient to only use a subset of the best models. The median skill of the 10 weighted members is better than using the full ensemble for all regions and sub-selection naturally leads to a highly reduced spread, which can be valuable for certain applications. However, to minimize the risk of making the projections worse it can be important to use the full weighted ensemble. This is particularly crucial for applications where worst-case scenarios are important, e.g. when the goal is to explore nonlinearities in impacts.

For precipitation the above considerations are even more important with the 5th percentile of the skill metric being reduced by 75% and more in many regions when picking only 10 models. For Sibiu and Madrid selecting only 10 models even leads to negative median skill (figure S5(b)). Using the full weighted ensemble does also not lead to significant skill increases, but at least limits the maximum CRPS decrease to about 50%.

In general, the model-as-truth tests show that weighting precipitation projections only leads to increased skill in reasonably large regions. In the largest region, the combined European domain hardly any increase in the median skill is found, which can be interpreted as the region being too large to apply a single weight per model. The highest skill increases are achieved in the Northern European region, with a median increase of about 40%.

4. Discussion, conclusions, and outlook

This work applies a weighting approach, which accounts for performance and independence, to temperature and precipitation projections in Europe. We show that weighting generally leads to an increase in the median projection skill, which can reach up to 50%. This result depends strongly on the region size, with smaller regions showing less increase and even a decrease in one extreme case. It is, therefore, evident, that some degree of spatial aggregation is needed in order to produce robust weighting. In turn, weighting leads to temperature and precipitation changes with opposing signs in northern and southern Europe, which cancel each other out when aggregating the full European region. Choosing the right region size is therefore crucial for producing reliable and physically meaningful results.

The European SREX regions are found to be such regions. For temperature the strongest effect of weighting is found in the Mediterranean, with a warming of over 2 °C (compared to 1995–2014) becoming very likely already by the middle of the century. For precipitation a large reduction in spread is evident in the Northern European SREX region. Our results reveal that a future decrease in Northern European precipitation is more likely than indicated by the unweighted multi-model mean with models which show an increase consistently getting the lowest weights.

Several important questions are not addressed here and will be investigated in future work: How does the number of diagnostics affect the weights and therefore the skill? Is there a subset of diagnostics which have predictive skill for many different targets? What is the impact of the diagnostic period length and the lead time? Is there an influence of model resolution on the ideal region size or is it rather driven by the homogeneity of the region? Further, the effect of the independence weighting is not discussed in detail here. It

has been argued that using a measure for model interdependence is important in so-called ‘ensembles of opportunity’ such as CMIP5, which are not designed to represent independent realizations of the climate system (e.g. Abramowitz *et al* 2019). An in-depth investigation of the independence weighting, e.g. by consecutively adding more and more initial condition members of the same model, is also currently in progress. Accounting for larger initial condition ensembles will become increasingly important in light of the upcoming CMIP6, for which several models with 20 and more members are planned (WCRP 2019).

To inform the weighting, we propose a novel approach for including observational uncertainty and apply different methods to account for the observational spread. In general, using only one observational estimate as ground truth does not account for the fact that datasets derived from observations represent reality with uncertainty. In the case of *in situ* measurements uncertainty arises from measuring itself and the subsequent processing into a gridded product. Specific sources of uncertainties are, for instance, the accuracy of the sensor, data gaps, and the gridding and interpolation procedures used. To account for the parametric uncertainties of the underlying procedures increasingly multiple realizations are constructed, such as the E-OBS or HadCRUT4 ensembles (Morice *et al* 2012, Cornes *et al* 2018). Satellite remote sensing such as the radiation measurements from the CERES instruments provide another important source of information for climate science. To derive relevant atmospheric variables on a regular grid from the raw measurements, physical theory and complex modeling as well as calibration procedures are necessary, which are subject to uncertainty. Finally, reanalysis datasets use process-based numerical models in their estimation and therefore additional uncertainties arise, for example, initial condition uncertainty and boundary condition uncertainty.

To account for these uncertainties we use three structurally different observational estimates and apply different methods to combine them. Our results show that including observational uncertainty leads to more robust results, mostly independent of the used method and therefore underline the importance of using multiple observational datasets. However, while the proposed weighting approach takes into consideration that climate models are not independent from each other, we still imply ‘observational democracy’, assuming that all observational estimates are independently constructed. Further research is needed to account also for inter-dependence between observational estimates which could then potentially allow the combined use of parametric and structural ensembles.

The selection of variables, regions, and time periods in this work is guided by a coordinated effort to compare different methods to constrain future projections of European climate, which is carried out in the

frame of the EUCP project (Hewitt and Lowe 2018). In light of the upcoming sixth assessment report (AR6) of the IPCC and with the new generation of CMIP6 models becoming available and the connected increase in initial condition members a broad discourse about ways to combine multi-model ensembles will become even more important in order to produce the best possible estimates of future changes.

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Data availability statement

The data that support the findings of this study are openly available at the links listed below. ERA-Interim is available at <https://apps.ecmwf.int/datasets/data/interim-full-mod/>, MERRA2 at <https://gmao.gsfc.nasa.gov/reanalysis/MERRA-2>, the CERES EBAF-Surface Product at <https://ceres.larc.nasa.gov/products.php?product=EBAF-Surface>, E-OBS at <https://ecad.eu>, and CMIP data at <https://esgf.llnl.gov/>.

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References

- Abramowitz G and Bishop C H 2015 Climate model dependence and the ensemble dependence transformation of CMIP projections *J. Clim.* **28** 2332–48
- Abramowitz G *et al* 2019 ESD Reviews: model dependence in multi-model climate ensembles: weighting, sub-selection and out-of-sample testing *Earth Syst. Dyn.* **10** 91–105
- Annan J D and Hargreaves J C 2017 On the meaning of independence in climate science *Earth Syst. Dyn.* **8** 211–24
- Baumberger C, Knutti R and Hirsch Hadorn G 2017 Building confidence in climate model projections: an analysis of inferences from fit *Wiley Interdiscip. Rev. Clim. Change* **8** 1–20
- Coppola E, Giorgi F, Rauscher S and Pianì C 2010 Model weighting based on mesoscale structures in precipitation and temperature in an ensemble of regional climate models *Clim. Res.* **44** 121–34
- Cornes R C, van der Schrier G, van den Besselaar E J and Jones P D 2018 An ensemble version of the E-OBS temperature and precipitation datasets *J. Geophys. Res. Atmos.* **123** 9391–409
- Cox P M, Huntingford C and Williamson M S 2018 Emergent constraint on equilibrium climate sensitivity from global temperature variability *Nature* **553** 319–22
- Dee D P *et al* 2011 The ERA-Interim reanalysis: configuration and performance of the data assimilation system *Q. J. R. Meteorol. Soc.* **137** 553–97
- Deser C, Knutti R, Solomon S and Phillips A S 2012a Communication of the role of natural variability in future North American climate *Nat. Clim. Change* **2** 775–9
- Deser C, Phillips A S, Bourdette V and Teng H 2012b Uncertainty in climate change projections: the role of internal variability *Clim. Dyn.* **38** 527–46
- Evans J P, Ji F, Abramowitz G and Ekström M 2013 Optimally choosing small ensemble members to produce robust climate simulations *Environ. Res. Lett.* **8** 044050
- Eyring V, Bony S, Meehl G A, Senior C A, Stevens B, Stouffer R J and Taylor K E 2016 Overview of the Coupled Model Intercomparison Project Phase 6 (CMIP6) experimental design and organization *Geosci. Model Dev.* **9** 1937–58
- Eyring V *et al* 2019 Taking climate model evaluation to the next level *Nat. Clim. Change* **9** 102–10
- Fatichi S *et al* 2016 Uncertainty partition challenges the predictability of vital details of climate change *Earth's Future* **4** 240–51
- Gelaro R *et al* 2017 The modern-era retrospective analysis for research and applications, version 2 (MERRA-2) *J. Clim.* **30** 5419–54
- Gleckler P J, Taylor K E and Doutriaux C 2008 Performance metrics for climate models *J. Geophys. Res. Atmos.* **113** 1–20
- Gordon N D and Klein S A 2014 Journal of geophysical research: atmospheres Low-cloud optical depth feedback in climate models *J. Geophys. Res. Atmos.* **119** 6052–65
- Hall A, Cox P, Huntingford C and Klein S 2019 Progressing emergent constraints on future climate change *Nat. Clim. Change* **9** 269–78
- Hall A and Qu X 2006 Using the current seasonal cycle to constrain snow albedo feedback in future climate change *Geophys. Res. Lett.* **33** 1–4
- Hawkins E and Sutton R 2009 The potential to narrow uncertainty in regional climate predictions *Bull. Am. Meteorol. Soc.* **90** 1095–108

- Hegerl G C, Brönnimann S, Schurer A and Cowan T 2018 The early 20th century warming: anomalies, causes, and consequences *Wiley Interdiscip. Rev. Clim. Chang.* **9** 1–20
- Herger N, Abramowitz G, Sherwood S, Knutti R, Angéllil O and Sisson S A 2019 Ensemble optimisation, multiple constraints and overconfidence: a case study with future Australian precipitation change *Clim. Dyn.* **00**
- Hersbach H 2000 Decomposition of the continuous ranked probability score for ensemble prediction systems *Weather Forecast.* **15** 559–70
- Hewitt C D and Lowe J A 2018 Toward a European climate prediction system *Bull. Am. Meteorol. Soc.* **99** 1997–2001
- IPCC 2012 *Managing the Risks of Extreme Events and Disasters to Advance Climate Change Adaptation. A Special Report of Working Groups I and II of the Intergovernmental Panel on Climate Change* (Cambridge: Cambridge University Press)
- IPCC 2013 *Climate Change 2013: The Physical Science Basis. Contribution of Working Group I to the Fifth Assessment Report of the Intergovernmental Panel on Climate Change* (Cambridge: Cambridge University Press)
- Kato S, Loeb N G, Rose F G, Doelling D R, Rutan D A, Caldwell T E, Yu L and Weller R A 2013 Surface irradiances consistent with CERES-Derived top-of-atmosphere shortwave and longwave irradiances *J. Clim.* **26** 2719–40
- Knutti R 2010 The end of model democracy? *Clim. Change* **102** 395–404
- Knutti R, Masson D and Gettelman A 2013 Climate model genealogy: generation CMIP5 and how we got there *Geophys. Res. Lett.* **40** 1194–9
- Knutti R, Sedláček J, Sanderson B M, Lorenz R, Fischer E M and Eyring V 2017 A climate model projection weighting scheme accounting for performance and interdependence *Geophys. Res. Lett.* **44** 1909–18
- Leduc M, Laprise R, de Elía R and Šeparović L 2016 Is institutional democracy a good proxy for model independence? *J. Clim.* **29** 8301–16
- Lorenz R, Herger N, Sedláček J, Eyring V, Fischer E M and Knutti R 2018 Prospects and caveats of weighting climate models for summer maximum temperature projections over North America *J. Geophys. Res. Atmos.* **123** 4509–26
- Morice C P, Kennedy J J, Rayner N A and Jones P D 2012 Quantifying uncertainties in global and regional temperature change using an ensemble of observational estimates: the HadCRUT4 data set *J. Geophys. Res. Atmos.* **117** 1–22
- O'Neill B C, Kriegler E, Riahi K, Ebi K L, Hallegatte S, Carter T R, Mathur R and van Vuuren D P 2014 A new scenario framework for climate change research: the concept of shared socioeconomic pathways *Clim. Change* **122** 387–400
- Parker W S 2009 II—Confirmation and adequacy-for-purpose in climate modelling *Proc. Aristot. Soc. Suppl. Vol.* **83** 233–49
- Pennell C and Reichler T 2011 On the effective number of climate models *J. Clim.* **24** 2358–67
- Sanderson B M and Knutti R 2012 On the interpretation of constrained climate model ensembles *Geophys. Res. Lett.* **39** 1–6
- Sanderson B M, Knutti R and Caldwell P 2015a A representative democracy to reduce interdependency in a multimodel ensemble *J. Clim.* **28** 5171–94
- Sanderson B M, Knutti R and Caldwell P 2015b Addressing interdependency in a multimodel ensemble by interpolation of model properties *J. Clim.* **28** 5150–70
- Sanderson B M, Wehner M and Knutti R 2017 Skill and independence weighting for multi-model assessments *Geosci. Model Dev.* **10** 2379–95
- Seneviratne S I *et al* 2018 The many possible climates from the Paris Agreement's aim of 1.5 °C warming *Nature* **558** 41–9
- Stegehuis A I, Teuling A J, Ciais P, Vautard R and Jung M 2013 Future European temperature change uncertainties reduced by using land heat flux observations *Geophys. Res. Lett.* **40** 2242–5
- Taylor K E, Stouffer R J and Meehl G A 2012 An overview of CMIP5 and the experiment design *Bull. Am. Meteorol. Soc.* **93** 485–98
- Tebaldi C and Knutti R 2007 The use of the multi-model ensemble in probabilistic climate projections *Philos. Trans. R. Soc. A* **365** 2053–75
- United Nations 2015 Paris Agreement (Accessed: 30 April 2019: http://unfccc.int/files/essential_background/convention/application/pdf/english_paris_agreement.pdf)
- van Vuuren D P *et al* 2011 The representative concentration pathways: an overview *Clim. Change* **109** 5–31
- Vogel M M, Zscheischler J and Seneviratne S I 2018 Varying soil moisture-atmosphere feedbacks explain divergent temperature extremes and precipitation projections in central Europe *Earth Syst. Dyn.* **9** 1107–25
- WCRP 2019 World Climate Research Programme CMIP Phase 6 (CMIP6) homepage (Accessed: 8 May 2019: <https://wcrp-climate.org/wgcm-cmip/wgcm-cmip6>)
- WMO 2019 WMO Statement on the State of the Global Climate in 2018. WMO-No.1233 (https://library.wmo.int/doc_num.php?explnum_id=5789)
- Zhang B and Soden B J 2019 Constraining climate model projections of regional precipitation change *Geophys. Res. Lett.* **2019** GL083926